Duration of Administration of Antibiotic Agents for Open Fractures: Meta-Analysis of the Existing Evidence

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Abstract

Background: Surgical site infection remains a significant concern in treating patients with open fractures. In prevention of such, current guidelines support the immediate administration of antibiotic agents. The duration of antibiotic treatment is still controversial. A maximum of 72 hours, even in the absence of definitive soft tissue coverage, is recommended in a number of recent guidelines and consensus reports. The aim of this meta-analysis was to review and analyze all published literature evidence with regard to antibiotic duration in open fracture treatment. Methods: We conducted a comprehensive review of the available literature from the 1970s until the present, including five comparative (1284 open fractures) and 27 observational (5408 open fractures) studies. A subgroup analysis was further performed, based on the Gustilo type of open injury and the anatomic location of the fracture. In addition, we investigated the effect of antibiotic regimes shorter than 72 hours on infection rates. Results: In the comparative studies, the summarized estimate of infection rate favored less than a 72-hour duration of antibiotic treatment, because prolongation of antibiotic treatment more than 72 hours did not seem to offer any protective effect against septic complications of open fractures (odds ratio: 0.85, 95% confidence interval [CI]: 0.60–1.21). The same trend was also documented in the observational studies, where the overall pooled estimate of infection rate was 10% (95% CI: 6.8%-14%) when antibiotic treatment did not exceed 72 hours and 9.2% (95% CI: 6.6%–12.2%) for more than 72 hours of antibiotic treatment (p=0.53). In Gustilo I and II open fractures, the calculated pooled estimate of infection rate did not differ significantly when antibiotic treatment exceeded 72 hours (6%, 95% CI: 3.3%–9%) compared with shorter (up to 72 h) antibiotic protocols (4%, 95% CI: 1.8%–7%) (p=0.52). In Gustilo III open fractures also, the calculated pooled estimate of infection rate (21.3%, 95% CI: 13%–31%) when duration of antibiotic treatment was more than 72 hours did not differ significantly compared with a shorter (less than 72 h) antibiotic treatment (17.7%, 95% CI: 12.5%-23.5%) (p=0.39). A further subgroup analysis indicated that even shorter antibiotic regimes (24–48 h) were also equivalent to prolonged regimes of more than 72 hours in terms of infection rates.

Conclusions: The results of the present systematic review and meta-analysis could not substantiate any benefit against septic complications of a prolonged duration of antibiotic treatment of open fractures, irrespective of their severity.

Keywords: antibiotics; duration of treatment; meta-analysis; open fractures

INFECTION RATES OF OPEN FRACTURES remain high, particularly with Gustilo III-type fractures [1,2]. Nosocomial superinfection of open fractures is an increasing problem in the hospital environment [3]. Immediate "prophylactic" antibiotic agent administration for visibly or potentially contaminated open fractures has been shown to reduce infection rates [4,5]. Historically, various antibiotic regimes have been used and their efficiency published [6]. The choice of duration in the antibiotic treatment of patients with open fractures remains a clinical challenge, because 30 years onward, very little is known about the appropriate duration of the antibiotic agents.

In elective orthopedic practice, single-dose antibiotic agents are sufficient to prevent peri-operative infections, as has been shown in joint arthroplasties [7].

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The guidelines published by the British Orthopaedic Association [8], which strictly speaking refer to the management of tibial open fractures, advocate the use of broadspectrum antibiotic agents in the emergency department (or even earlier—i.e., on site or in the ambulance [5,9,10] and at the time of wound debridement to decontaminate the wound [Step 1: decontamination]) and recommend the administration of teicoplanin (or gentamicin) at induction at the time of skeletal stabilization and definitive soft tissue cover to reduce the risk of hospital-acquired infections (Step 2: prevention of nosocomial infection). Co-amixoclav should be administered concomitantly until definitive wound cover has been established or for a maximum of 72 hours. These measures in combination with timely debridement, irrigation, and early definitive soft tissue cover by an orthoplastic team are expected to minimize the risk of post-injury infection [11].

The aim of the present meta-analysis was to validate the consensus opinion that short-term antibiotic protocols (duration of intravenous antibiotic treatment up to 72 h or less) are as efficient as longer-term regimes in reducing infection rates of all long bone fractures.

Methods

Literature search and data extraction

We conducted a systematic and comprehensive review of the existing literature adhering to the PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines [12]. Before commencing the comprehensive literature search, a written study protocol was established including clearly defined eligibility criteria and specifying the criteria for further subgroup and sensitivity analyses. Clinical studies dealing with open long bone fractures with a clear documentation of the duration of the utilized antibiotic protocol and reporting on infection rates in association with the duration of the respective antibiotic protocol were regarded

		Section	Number or factor	Score
Part A	1	Study size (number of	>60	10
		patients)	41-60	7
		•	20–40	4
			<20, not stated	0
	2	Mean follow-up (months)	>24	5
			12–24	2
			<12, not stated or unclear	0
	3	No. of different antibiotic	One antibiotic protocol only	10
		protocols included in each reported outcome. More	> 1 protocol, but >90% of subjects undergoing this protocol	1
		than one antibiotic protocol may be assessed, but separate outcomes should be reported	Not stated, unclear, or <90% of subjects undergoing the one protocol	0
	4	Type of study	RCT	15
		Type of study	Prospective cohort study	10
			Retrospective cohort study	0
	5	Diagnostic certainty	In all	5
		8	>80%	3
			<80%, not stated or unclear	0
	6	Description of surgical	Adequate	5
		protocol	Fair	3
	_		Inadequate, not stated or unclear	0
	7	Description of antibiotic	Well described	10
		protocol	Not adequately described	5
			Protocol not reported	0
Part B	1	Outcome criteria	Outcome measures clearly defined	2
			Timing of outcome assessment clearly stated	2
			Use of outcome with good sensitivity	3
			Use of outcome criteria that has reported good reliability	3
	2	Procedure for assessing outcomes	Subjects recruited (results not taken from surgeons' files)	8
			Investigator independent of surgeon/therapist	7
	3	Description of subject	Selection criteria reported and unbiased	5
		selection process	Recruitment rate reported >80% or recruitment rate reported <80%	5 or 3
			Eligible subjects not included in the study satisfactorily accounted for (e.g., dropout analysis), or 100% recruitment	5
Maximu	n score		or room recruitment	100

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RCT = randomized controlled trial.

DURATION OF ANTIBIOTICS FOR OPEN FRACTURES

eligible for inclusion in the systematic review. Exclusion criteria included pediatric open fractures, open spinal injuries, open fractures of hand, foot, and maxillofacial region, experimental studies, animal studies, case reports, and studies containing fewer than 20 subjects.

A comprehensive electronic search of MedLine via the PubMed search machine was performed using the following search terms and Boolean operators: "open fracture" AND "antibiotics." The search was further extended to the Ovid MEDLINE, CINAHL, Cochrane Library, Embase, Google Scholar, and Scopus databases. No language restrictions were set. Further, other relevant publications (such as reviews and meta-analyses) were obtained, and their bibliographies were searched manually for potentially eligible papers. Any disagreement between the two reviewers was resolved by discussion. From each eligible article, information on author's name, year of publication, type of study, demographic and baseline characteristics of participants, follow-up details, outcome data, and complications were extracted and documented on an Excel sheet. No limitation was set a priori with respect to the type of eligible studies, although our priority was to include high quality comparative studies.

Quality assessment

The methodologic quality of all primary studies was evaluated with the Coleman Methodology score (CMS) [13]. The total score can range from 0 to 100, and higher scores are indicative generally of absence of various biases and confounding factors. The final score was categorized as excellent

Statistical analysis

For comparative studies, pooling of data was performed with the Mantel-Haenszel (M-H) statistical method and an either fixed or random effects model, depending on the degree of the statistical heterogeneity present (in the presence of significant statistical heterogeneity, a random effects model was used). Binary outcomes were summarized as odds ratios (ORs) with 95% confidence intervals (95% CI). The results of each primary study and the combined estimate of effect size were presented graphically as forest plots. Statistical heterogeneity was measured with the use of both Cochran's X^2 (Q-test) and I^2 statistics [14,15]. Significance was set at 0.1 for the Q test (because it is characterized by low sensitivity for detecting heterogeneity). An I^2 value >50% was thought to represent significant heterogeneity. The RevMan (5.2) software (Review Manager, The Nordic Cochrane Centre, Copenhagen, Denmark) was used to present the study findings, produce pooled estimates of effect size, and test the presence of statistical heterogeneity.

PRISMA Flow Diagram Open fractures AND antibiotics



FIG. 1. Preferred Reporting Items for Systematic Reviews and Meta-Analyses flowchart.

			TABI	le 2. Summ	IARY OF A	rticles Included in this Meta-Analysi	sis Review		
First author (Ref)	Year	Level	No of pts*	: Site	Grade	Antibiotic protocol	Duration	CMS	Wound/surgical management/comment
Dellinger EP, et al. (18)	1988	2b	248	All	Compar I–III	ative studies (<72 h vs. >72 h of antibiotic Cefonicid (24 h), cefonicid (5 d),	administration, n= 24 h vs. 5 d	=5 66	
Merritt K (19)	1988	4	68	All	III-I	cetamandole (3 d) Cephalosporin±aminoglycoside	72h vs. 5d	67	ORIF $(n=35)$, cast $(n=16)$, $r_{22} = r_{22} = r_{10}$
Carsenti-Etesse H,	1999	1a	616	Tibia	II –I	Pefloxacin single dose vs. oxacillin +	Single dose	87	EX-IIX (n = 19) Positive wound cultures did not
et al. (17) Barton CA, et al. (16) Rodriguez L, et al. (20)	2012 2014	44	214 174	All Tibia, femur		cetazoline Cefazolin Cefazolin (I-II), ceftriaxone + aminoglycoside glycopeptide (III)	vs. >/2h <72h vs. >72h 72h vs. >72h	47 48	predict intective organism
Clancey CJ, et al. (26)	1978	4	98	<i>Ob</i> Tibia	servationa I-III	l non-comparative studies (<72 h of antibi Methicillin + kanamycin	otic administration 72 h	n, n = 53	14 Cast treatment $(n = 56)$, IMN (n = 13), ORIF $(n = 22)$, Ex-fix
Bergman BR (23) Gustilo RB, et al. (31) Franklin II., et al. (30)	$1982 \\ 1984 \\ 1984$	4 2b 2b	60 207 42	Ankle All Ankle	III-1 III-1	Dicloxacillin or benzylpenicillin Cephalosporin Cenhalosnorin	48 h 72 h 48 h	51 60	(n = /) Predominantly Ex-fix treatment Delayed closure and ORIF All delayed closure at 5 d. 8
Johnson KD, et al. (33)	1988	1c	46	Tibia	III-I	Cefazolin + cefotaxime	48 h	58	SSGs $Ex-fix (n=37)$, IMN $(n=5)$, cast
Patzakis MJ, et al. (38) Bednar DA, et al. (21)	1989 1993	4 2b	109 82	All	111-1 111-1	Cefamandole + tobramycin Cephalosporin±tobramycin	72 h 48 h	41 67	(n = 3), OKLF $(n = 1)ORIF (n = 45), intramedullary$
Cole JD, et al. (27)	1995	4	50	Tibia	III-I	Cefazolin±aminoglycoside (+ local bead	ls 36h	39	nails (n=29), Ex-fix (n=8)
Templeman DC, et al. (43)	1998	4	133	Tibia	III-I	in GIII n = 3) Cefazolin (I), cephalosporin + aminoglycoside (II-III),	72 h	37	Unreamed IMN $(n=67)$, reamed IMN $(n=3 2)$, cast $(n=14)$
Vasenius J, et al. (44)	1998	1c	240	All	III-I	later ticarcillin (+tobramycin beads) Clindamycin + cloxacillin vs.	72 h	61	ORIF (n=136), Ex-fix (n=42)
DeLong WG, et al. (28) Dunkel B, et al. (29) Leonidou A, et al. (35)	$1999 \\ 2013 \\ 2014$	444	118 1492 68	All All All		Cefazolin Cefuroxime Co-amoxiclav	72 h 72 h 72 h	52 42 74	68 treated according to BOA
Lack WD, et al. (5)	2015	1c	137	Tibia	III-I	Cefazolin	48 h	49	guidelines 125 surgically debrided within 24h
Christensen J, et al. (25)	1982	4	40	<i>Obser</i> Tibia	vational n I- II	on-comparative studies (>72 hours of antil Methicillin + gentamycin IV	ibiotic administrati 12 d	ion, n 41	= <i>13</i>) Immediate debridement and

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(continued)

						TABLE 2. (CONTINUED)			
First author (Ref)	Year	Level	No of pts,	* Site	Grade	Antibiotic protocol	Duration	CMS	Wound/surgical management/comment
Patzakis MJ, et al. (40)	1983	4	109	Tibia	III-1	Cefamandole + tobramycin	3–5 d	69	fixation If cultures negative 3 days, if positive 5 days of antibiotic treatment; cast treatment (n = 69) (25 cast + pin), Ex-fix (n = 15), primary closure (n = 19), partial $(n = 75)$, left
Benson DR, et al. (22)	1983	2b	82	All	III-I	Cefazolin vs. clindamycin	5 d	73	open $(n = 1.5)$. No difference between study
Braun R, et al. (24)	1987	2b	47	All	III-I	Cloxacillin vs. placebo	10d	88	groups 86/91 primary fixation, 70 plates, 5 TBW, 1 IMN, 6 screw
Patzakis MJ, et al. (38)	1989	4	1104	ИI	III-I	Penicillin + streptomycin or cephalothin or	3-10 d	39	Various antibiotic regimens compared
Russell GG, et al. (41)	1990	4	37	Tibia	III-I	ceramancole + looramycin Cephalosporin or penicillin + cloxacillin	>72 h	27	Cast $(n = 63)$, ORIF $(n = 7)$, Ex. 62, (2 - 7)
Seligson D, et al. (42)	1994	1b	32	ЯII		Penicillin, cefazolin±tobramycin beads	5 d	53	Significantly $(u = v)$ in tobramycin beads group in tobramycin beads group
Osterman PA, et al (37) Vasenius J, et al (43)	1995 1998	4 1c	240 109	All All	III-I III-I	Penicillin or tobramycin or cefazolin Cloxacillin	>72 h 4 d	50 61	Systemic vs. local administration Wound swab, then 1L saline,
Moehring HD, et al (36)	2000	1c	38	All	qIII-II	Cefazolin±aminoglycoside (IIIA/ B)±Metronidazole (contaminated) as single dose in ER, cephalosporin IV post_admission	>96h	70	Saline irrigation in ER, definitive treatment within 6 h, IMN $(n = 38)$
Gopal S, et al (31)	2000	4	84	Tibia	Ш	Cefuroxime + metronidazole	5 d	46	Immediate "fix and flap"
Patzakis MJ, et al (39)	2000	1b	171	All	III-I	Ciprofloxacin or cefamandole +	3–8 d	64	approach Debridement within 24 h, 10 L soline irrigation
Lenarz CJ, et al (34)	2010	4	422	All	III-II	Cephalosporins + aminoglycoside	3 wks average	60	Repeat irrigation and debridement until wound cultures are negative
			4						

pts = patients; CMS = Coleman Methodology Score; ORIF = open reduction internal fixation; Ex-fx = external fixation; IMN = intra-medullary nailing; SSG = split skin graft; BOA = British Orthopaedic Association; IV = intravenous; TBW = tension band wiring; ER = emergency room. *Number of patients/firactures relevant to our study question (not necessarily reflecting the whole study population in the publication).



FIG. 2. Funnel plot of the results reported to the comparative studies that were analyzed in this meta-analysis. Abs = antibiotic agents; CI = confidence interval.

TABLE 3. COLEMAN METHODOLOGY SCORE: MEAN, MEDIUM, STANDARD DEVIATION, AND RANGE VALUES FOR ALL COMPONENT STUDIES AND FOR COMPARATIVE AND NON-COMPARATIVE STUDIES, SEPARATELY

		CN	AS score	
Type of studies	Mean	SD	Range	Median
All studies (5, 15–43) Comparative studies (15–19) Non-comparative studies (5, 20–43)	56 63 54	14 16 14	27–88 47–85 27–88	53 66 53

CMS = Coleman Methodology Score; SD = standard deviation.

For observational studies without a comparator cohort, all outcomes of interest were expressed as proportions (p), (infection rates). Pooling of proportions was performed with the MedCalc software (version 14.8.1) using a random effects model. Statistical heterogeneity was also tested with Cochran's Q test and Higgins I^2 test. Non-parametric comparisons of the median values of outcomes of interest between two groups were performed with the Mann Whitney U test.

Subgroup analysis

The type of open fracture (according to the Gustilo classification) and the anatomic site of the open fractures were the a priori set criteria for subgroup analysis. After reviewing the primary studies, we realized that there were even shorter antibiotic prophylactic regimes of up to 24 or 48 hours that were used by some authors. Consequently, we also compared those regimes with the longer antibiotic prophylaxis of more than 72 hours in terms of infection rate.

Sensitivity analysis

The criteria of sensitivity analysis were also pre-specified and included studies of dubious eligibility, poor methodologic quality, or outlying results. We intended to repeat the analysis after excluding studies fitting the above categories. We would regard the results of our review with greater certainty if the process of sensitivity analysis did not affect them.

Results

Search process

The initial electronic search generated 847 results. Another 53 records were identified through published guidelines and consensus documents. After duplicates were removed, 794 abstracts and abstract titles were screened for suitability. For full article review, 57 publications were retrieved and analyzed. After applying eligibility criteria, 32 studies were left for final analysis (see PRISMA flowchart, Fig. 1). Five of them were comparative studies [16–20], while the remaining 26 were observational non-comparative studies [5,19,21–44]. In one comparative study [18], we analyzed separately two different treatment groups, based on the type of the used antibiotic prophylaxis. Subsequently, there were six pairs of treatment groups available for pooled analysis out of five comparative studies.

Data were extracted regarding study design, patient numbers, patient demographics, antibiotic and surgical protocol, infection and complication rates. If different antibiotic duration regimens were used within a study, we split the patient populations and analyzed them separately. Both superficial and deep infections were extracted as one category. Table 2 shows a summary of all studies included, their CMS score, the various antibiotic regimes followed, the duration of administration, the study population, and the level of evidence.

Publication bias

We generated a funnel plot of infection rate for all primary comparative studies (used in the meta-analysis) to assess the likelihood of publication bias. The distribution of data points



FIG. 3. Foster plot of the results reported to the comparative studies that were analyzed in this meta-analysis. Abs = antibiotic agents; CI = confidence interval.

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Table 4. Comparison of Non-Comparative Studies as to the Median Incidence of Infection Rates of Open Fractures Managed with Protocols of Antibiotic Administration for Either More Than 72 Hours or Less Than 72 Hours

	Munhar	Infaction	Medica value	Woidhted mean		Heterogen	eity
Authors; year; reference	of cases	rate	of infection rate	weigned mean of infection rate	95% CI	Q (df)	I^{2} (95%CI)
Duration of antibiotic agents >72 h Christensen J, et al.; 1982; (25) Patzakis MJ. et al.: 1983; (40)	40 109	0 4.6%	6.4%*	9.2% ^a	6.6%-12.2%	Q=68 (14), p<0.001	$I^2 = 80 \ (67 - 87)$
Benson DR, et al.; 1983; (22)	82	6.1%					
Braun R, et al.; 1987; (24) Dotrobie MT at al. (1980, (23) (Don ± Strent)	47 210	4.3%					
Patzakis MJ, et al.; 1989; (38) (Cef)	210	5.5%					
Russell GG, et al.; 1990; (41)	90 2	14.4%					
Seligson D, et al.; 1994; (42) Octarman DA at al · 1005· (37)	25 270	20 % 10 10%					
Vasenius J. et al.: 1998: (44)	109	20.2%					
Mochring HD, et al., 1999; (36)	38	5.3%					
Patzakis MJ, et al.; 2000; (39) (ciprofloxacin) Patzakis MJ, et al.; 2000; (39) (Cef + Gentam)	78 93	14.1% 6.45%					
Gopal S, et al.; 2000; (31) Lenarz CJ, et al.; 2010; (34)	84 422	15.5% 4.3%					
Duration of antihiotic agents <77 h							
Clancey GJ, et al.; 1978; (26)	98	15.3%	10%*	$10 \ \%^{a}$	$6.8\%{-}14\%$	Q=96.6 (13), p<0.001	$I^2 = 86.5 \ (79-91)$
Bergman BR; 1982; (23)	60	6.7%					
Franklin JL, et al.; 1984; (30)	42	16.7%					
Gustilo RB, et al.; 1984; (32)	207	10.6%					
Johnson KD, et al.; 1988; (33)	40 0 0	21.8%					
Bednar DA, et al.; 1993; (21)	2.8 8.7	4.9%					
Cole JD, et al.; 1995; (2/) Templeman DC et al · 1008· (43)	00 133	$\frac{2}{1130}$					
Vacentine I et al · 1008· (44) (Clind ve clov)	CC1	14.50%					
Vascuus J. et al.: 1998: (44) (Clind)	117	9.3%					
DeLong WG Jr, et al; 1999; (28)	118	6.8%					
Dunkel N, et al.; 2013; (29)	1492	3.6%					
Leonidou A, et al; 2014; (35)	68	7.3%					
Lack WD, et al; 2015; (5)	137	17.5%					

CI = confidence interval; df = degrees of freedom. *p=0.53 (Mann-Whitney U test). *Random effects model.



FIG. 4. Foster plot per Gustilo classification groups of the results reported in comparative studies that were analyzed in this meta-analysis. Abs=antibiotic agents; CI=confidence interval.

within the funnel plot was almost symmetric, implying that presence of publication bias was unlikely (Fig. 2).

Quality assessment

The CMS ranged from 27 to 88 across all primary studies (mean: 56, standard deviation: 14, median: 53). Comparative studies scored a higher score compared with non-comparative ones (Table 3). The ICC was 0.92 (95% CI: 0.81–0.99), implying a nearly perfect agreement between the two assessors.

Infection rates

1. Comparative studies. Five studies (six treatment groups) [16–20] provided relevant data. The pooled estimate of effect size for infection rate did not document any statistically significant difference between the two groups, although it seemed to favor a protocol of less than 72 hours duration of antibiotic agents (OR: 0.85, 95% CI: 0.60–1.21) in the absence of statistical heterogeneity (Q=4.23, df: 5, p=0.37, $I^2=0$) (Fig. 3).

2. Observational studies (no comparator group). These studies were grouped accordingly, based on the duration of antibiotic treatment of their participants either more or less than 72 hours. The group of "duration of abx >72 hrs" consisted of 15 patient populations derived from 13 studies [22,24,25,31,34,36–42,44] (Table 4). The infection rate ranged from 0%–20% (median: 6.4%) across component studies. The pooled estimate of infection rate was calculated (weighted mean of infection rate: 9.2%, 95% CI: 6.6%–12.2%), but it should be interpreted with caution, because of

the presence of significant statistical heterogeneity across the primary studies (Q=68, df: 14, p < 0.001, $I^2 = 80$).

The group of "duration of abx <72 hrs" included 14 patient populations from 13 studies [5,21,23,26–30,32,33,35,43,44]. The infection rate ranged from 2%–22% (median: 10%). The pooled analysis was deemed with significant statistical heterogeneity (Q=96.6, df: 13, p<0.001, I^2 =86.5) and, again, the respective results should be interpreted cautiously (weighted mean of infection rate: 10%, 95% CI: 6.8%–14%). The documented difference between the above groups was not statistically significant (p=0.53, Mann Whitney U test) (Table 4).

Subgroup analysis

Subgroup analysis was performed in terms of Gustilo type, anatomic site of the open fractures, and shorter (<48 h) antibiotic regimes.

1. Comparative studies. These studies provided data for subgroup analysis based on the Gustilo type of open fractures. For this purpose, we stratified studies reporting on Gustilo type I + II open fractures and on Gustilo III open fractures (with no limitations per anatomic site). The results of subgroup analysis of the comparative studies are presented in Figures 4 and 5. The results did not support any advantage of duration of antibiotic treatment over 72 hours when open fractures were stratified per Gustilo type.

2. Observational studies. We were able to stratify them according to both the anatomic site and Gustilo type of open fracture. Concerning the anatomic site, only articles reporting

lange of the second	ABs duration <	72 hrs	ABs duration >	72 hrs		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI Ye	ar M-H, Fixed, 95% Cl
Dellinger (cefamandole)	7	29	8	37	19.0%	1.15 [0.36, 3.66] 19	8
Merritt	3	30	10	38	28.3%	0.31 [0.08, 1.25] 19	8 =
Dellinger (cefonicid)	7	29	8	39	18.4%	1.23 [0.39, 3.90] 19	8 =
Barton	2	15	13	61	15.8%	0.57 [0.11, 2.84] 20	2
Rodriguez	5	18	11	37	18.5%	0.91 [0.26, 3.17] 20	4
Total (95% CI)		121		212	100.0%	0.79 [0.45, 1.39]	•
Total events	24		50				
Heterogeneity: Chi ² = 2.91	, df = 4 (P = 0.57);	$I^2 = 0\%$					
Test for overall effect: Z =	0.82 (P = 0.41)						Favours duration<72hrs Favours duration>72hrs

FIG. 5. Foster plot of the results reported for tibial open fractures in comparative studies that were analyzed in this metaanalysis. Abs = antibiotic agents; CI = confidence interval.

			Median	Weighted		Hetero	geneity
Authors; year; reference	Number of cases	Infection rate	value of infection rate	mean of infection rate	95% CI	Q(df)	<i>I</i> ² (95% <i>CI</i>)
Duration of antibiotic agents >72	h						
Christensen J, et al; 1982; (25)	40	0	9.5%*	$7.8\%^{\mathrm{a}}$	2.2%-16.6%	Q = 18 (3),	$I^2 = 83$
Patzakis MJ, et al; 1989; (38)	109	4.6%				p<0.01	(57–93)
Russell GG, et al; 1990; (41)	90	14.4%				1	. ,
Gopal S, et al; 2000; (31)	84	15.5%					
Duration of antibiotic agents <72	h						
Clancey GJ, et al; 1978; (26)	98	15.3%	15.3%	13.3 ^a	8.0%-19.8%	Q = 14 (4),	$I^2 = 71$
Cole JĎ, et al; 1995; (27)	50	2 %				p<0.01	(27 - 88)
Johnson KD, et al; 1988; (33)	46	21.7%				1	. ,
Templeman DC, et al; 1998; (43)	133	11.3%					
Lack WD, et al; 2015; (5)	137	17.5%					

TABLE 5. COMPARISON OF INFECTION RATE OF OPEN TIBIAL FRACTURES IN THE OBSERVATIONAL STUDIES, STRATIFIED PER DURATION OF ANTIBIOTIC ADMINISTRATION

CI = confidence interval; df = degrees of freedom.

*p=0.4, Wilcoxon-Rank Sum Test.

^aRandom effects model.

on open tibial fractures were available for subgroup analysis (Table 5). The results did not favor antibiotic agent administration in excess of 72 hours (p = 0.4, Mann Whitney U test). About the Gustilo type of open fractures: An appropriate subgroup analysis was feasible, and the respective results are depicted in Tables 6 and 7. The results did not support duration of antibiotic treatment in excess of 72 hours for both Gustilo I+II (p=0.52, Mann Whitney U test) and Gustilo III subgroups (p=0.39, Mann Whitney U test).

3. Shorter antibiotic prophylactic regimes. (i) Comparative studies. We directly compared regimes of "antibiotic duration less than 24 hours" with the longer regime (antibiotic duration more than 72 hours) for all open

TABLE 6.	INFECTION	RATES FOR	GUSTILO	Types I	and II	Open	FRACTURES	s:
DUR	ATION OF A	NTIBIOTIC A	Agents >7	72 Hours	s Versu	us <72	Hours	

	N7 1		Median	Weighted		Hete	rogeneity
Authors; year; reference	Number of cases	Infection rate	value of infection rate	mean of infection rate	95% CI	Q(df)	I ² (95% CI)
Duration of antibiotic agents >72 h							
Christensen J, et al.; 1982; (25)	40	0	5.8%*	$6\%^{\rm a}$	3.3%-9%	Q = 20 (8),	$I^2 = 60 (16 - 81)$
Patzakis MJ, et al.; 1983; (40)	94	2.1%				p = 0.01	
Braun R, et al.; 1987; (24)	34	2.9%					
Russell GG, et al.; 1990; (41)	72	8.3%					
Seligson D, et al.; 1994; (42)	32	25 %					
Osterman PA, et al.; 1995; (37)	138	5.8%					
Patzakis MJ, et al.; 2000; (39) (ciprofloxacin)	52	5.8%					
Patzakis MJ, et al.; 2000; (39) (Cef + Gentam)	67	6.0%					
Lenarz CJ, et al.; 2010; (34)	123	4.0%					
Duration of antibiotics <72 h							
Clancey GJ, et al.; 1978; (26)	87	10.3%	2.0%*	$4\%^{\mathrm{a}}$	1.8%-7%	Q = 32 (9),	$I^2 = 72 (47 - 85)$
Franklin JL, et al.; 1984; (30)	26	19.2%				p<0.01	
Gustilo RB, et al.; 1984; (32)	125	1.6%				1	
Johnson KD, et al.; 1988; (33)	19	10.5%					
Bednar DA, et al.; 1993; (21)	19	$0 \ \%$					
Cole JD, et al.; 1995; (27)	19	$0 \ \%$					
Templeman DC, et al.; 1998; (43)	65	1.5%					
Vasenius J, et al.; 1998; (44) (clind)	87	2.3%					
DeLong WG Jr, et al.; 1999; (28) Dunkel N, et al.; 2013; (29)	68 1033	3.0% 1.3%					

df=degrees of freedom; CI=confidence interval.

*p=0.52 (Mann-Whitney U test). ^aRandom effects model.

	N7	I. C. dian	Median	Weighted		Hetero	geneity
Authors; year; reference	of cases	rate	infection rate	mean of infection rate	95% CI	Q(df)	I ² (95% CI)
Duration of antibiotic agents >72 h							
Patzakis MJ, et al; 1983; (40)	15	20 %	20.3%*	21.3% ^{&}	13 %-31%	Q = 50 (9),	$I^2 = 82$
Braun R et al; 1987; (24)	13	7.7%				P<0.001	(68–90)
Russell GG, et al; 1990; (41)	18	39 %					
Seligson D et al; 1994; (42)	32	25 %					
Osterman PA, et al; 1995; (37)	102	20.6%					
Vasenius J, et al; 1998; (44)	27	52 %					
Patzakis MJ, et al; 2000; (39) (ciprofloxacin)	26	31 %					
Patzakis MJ, et al; 2000; (39) (Cef + Gentam)	26	7.7%					
Gopal S. et al: 2000: (31)	84	15.5%					
Lenarz CJ, et al; 2010; (34)	173	5.8%					
Duration of antibiotic agents <72 h							
Clancey GJ, et al; 1978; (26)	11	54.5%	17.5%*	$17.7\%^{\rm a}$	12.5%-23.5%	Q = 35 (10),	$I^2 = 71$
Franklin JL, et al; 1984; (30)	16	12.5%				p<0.001	(47 - 84)
Gustilo RB, et al; 1984; (32)	82	24.4%					
Johnson KD, et al; 1988; (33)	27	29.6%					
Bednar DA, et al; 1993; (21)	63	6.3%					
Cole JD, et al; 1995; (27)	31	3.2%					
Templeman DC, et al; 1998; (43)	68	20.6%					
Vasenius J, et al; 1998; (44) (Clind)) 31	29 %					
DeLong WG Jr, et al; 1999; (28)	50	12 %					
Dunkel N, et al; 2013; (29)	310	12 %					
Lack WD, et al; 2015; (5)	137	17.5%					

TABLE 7. COMPARISON OF THE INFECTION RATES FOR GUSTILO TYPE III OPEN FRACTURES STRATIFIED PER DURATION OF ANTIBIOTIC AGENTS

df=degrees of freedom; CI=confidence interval.

*p=0.39 (Mann-Whitney U test),

^aRandom effects model

fractures (Table 8). We further stratified the results of the relevant studies based on the Gustilo type of open fractures (Fig. 6). The results did not show any clear benefit of prolonging antibiotic prophylaxis of open fractures even beyond the first 24 hours. (ii) Non-comparative studies. In this category of included studies, we were able to compare the regimes of "less than 48 hours antibiotic administration" with "more than 72 hours antibiotic administration." The respective results are depicted in Table 9. Again, no statistically significant difference could be established between the compared groups in terms of infection rate.

For sensitivity analysis, we excluded studies that had received fewer than 40 points according to the CMS because these were regarded as methodologically weaker [27,38,41,43]. Repeat pooling analysis did not produce materially different results compared with the original ones. We have also repeated the pooling analysis after first excluding studies with outlying results [33,37]. Again, this procedure did not produce substantially different results compared with the original ones.

Discussion

This is the first meta-analysis focusing specifically on the duration of antibiotic administration in all long bone open fractures. The results from pooled analysis of both comparative and observational studies do not substantiate that prolonged (more than 72 h) antibiotic schemes in open fractures offer any benefit against septic complications.

Open fractures posed a high death occurrence and morbidity to patients before the arrival of modern antibiotic agents. Injured soldiers in wars either faced immediate amputation of the affected limb or weeklong bed rest with

TABLE 8. SUBGROUP ANALYSIS OF COMPARATIVE STUDIES: POOLED ESTIMATES OF INFECTION RATES OF OPEN FRACTURES WITH RESPECT TO THE DURATION OF ANTIBIOTIC PROPHYLAXIS (LESS THAN 24 HOURS VERSUS MORE THAN 72 HOURS)

Outcome: Infection rate	Studies	Participants	Statistical method	Effect estimate
All open fractures, <24 vs. >72 h	3	943	Odds ratio (M-H, Fixed, 95% CI)	0.91 (0.58, 1.42)
Gustilo I & II, <24 vs. >48 h	3	808	Odds ratio (M-H, Fixed, 95% CI)	0.87 (0.50, 1.50)
Gustilo III, <24 s vs. >72 h	2	134	Odds ratio (M-H, Fixed, 95% CI)	1.19 (0.53, 2.70)

CI = confidence interval.



FIG. 6. Forest plot comparing antibiotic agent administration of fewer than 24 hours against more than 72 hours for all open fractures (**A**), Gustilo type I and II open fractures (**B**), and Gustilo type III open fractures (**C**). CI = confidence interval.

painful dressing changes. Both interventions were regularly complicated by severe infections, which most commonly led to the demise of the patient. The discovery of antibiotic agents and their use in open fractures reduced the infection rate significantly. The first study documenting their efficiency was published by Patzakis et al. in 1974 [45], which showed that cephalosporins reduced the infection rate by two thirds compared with placebo treatment. Ever since, various arbitrary antibiotic regimes have been used and their efficiency published in relatively small cohort studies. The time of administration, the combination of various agents, the duration of administration and their local or systemic application has been discussed widely in the literature [1].

Interestingly, almost 40 years after the publication by Patzakis et al., the duration of antibiotic treatment remains controversial. The choice of agent depends mainly on the community and hospital environment, but there is some agreement that a broadspectrum antibiotic should be used in the emergency department whereas more specific agents are necessary in primary and subsequent surgical procedures to avoid nosocomial multiresistant superinfection. The British Orthopaedic Association (BOA) guidelines published in 2009 [46] are based on a review article by Jaeger et al. [47], which takes previous meta-analyses and consensus publications into account. Since 2006, no further comprehensive meta-analysis has been undertaken, and this is the first meta-analysis and review focusing specifically on the duration of antibiotic administration.

We would like to suggest the elimination of the term "antibiotic prophylaxis" because it is being used in elective orthopedic practice. The nature and mechanism of an open fracture leads to a contamination of the wound site per definition, and the early administration of antibiotic agents in the emergency department has reduced the infection rate significantly [35,48,49]. This decontamination step does not need to be protracted over 72 hours, because the wound should have been debrided and irrigated [50,51], either primarily closed or covered or temporarily sealed within the first 24 hours 31,52]. Any prolongation of the initial antibiotic beyond 72 hours potentially increases the risk of noso-comial infection [53]. Vasenius et al. [44] found that resistance to the initial antibiotic was found in eight of 11 infected patients (clindamycin group) and 16 of 22 (cloxa-cillin group).

Many factors influence the outcome of open fracture treatment. Some paradigms are now outdated because more robust studies show that the old "six hour rule" does not apply [48]; similarly, wound swabs in the emergency department and at first surgical debridement are very poor predictors of the organism causing the subsequent infection [34]. Primary closure after thorough debridement, multi-disciplinary orthoplastic approach, and early coverage with soft tissue grafts [31] or negative pressure dressings have been shown to improve overall outcome [54].

Very little is still known about the appropriate management of heavily contaminated open fractures. There is a lack of studies specifically addressing farmyard and open water injuries. The choice of antibiotic agents is guided by local resistance profiles and varies highly among all studies.

Vasenius et al. [44] have shown that Grade III open fractures have gram-negative pathogens in 43% of cases. The BOA/ British Association of Plastic, Reconstructive and Aesthetic Surgeons (BAPRAS) guidelines recommend co-amoxiclav because it has a higher bacterial kill rate than cefuroxime and offers gram-positive, gram-negative, and anaerobic cover [8]. Teicoplanin offers excellent bone penetration, covers gram-positive

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				Infect	ion rate		
	Number of included patient populations (Refs)	Sample size	Range	Median	Weighted mean ^a (95% CI)	Heterogeneity	p*
<i>All data</i> ≤48 h ≥72 h	6 (5,20,22,26,29,32) 15 (21,23,24,30,33,35–40,43)	417 2380	2-21.8% 0-25%	$\frac{11.7\%}{6.4\%}$	$\frac{11\%}{9.2\%} (5.5\% - 18\%)$	Q = 21.3, df = 5, p = 0.0007, I^2 = 76.5% (95% CI: 47.5%-89.5%) Q = 68, df = 14, p < 0.0001, I ² : 79% (95% CI: 67%-87%)	0.57
<i>Open fra</i> ≤48 h ≥72 h	ctures, Gustilo I and II 4 (20,26,29,32) 9 (23,24,33,36,38–41)	83 652	$\begin{array}{c} 0-19\% \\ 0-25\% \end{array}$	5% 6%	7% (1%–19%) 6% (3.3%–9%)	Q=8.8, df=3, p=0.0321, I ² : 66% (95% CI: 0–88.4%) Q=20, df=8, p=0.011, I ² : 60% (95% CI: 16%–81%)	0.94
<i>Open fra</i> ≤48 h ≥72 h	ctures, Gustilo III 5 (5,20,26,29,32) 10 (23,30,33,36,38–41,43)	274 516	3.2–30% 5.8–52%	12.5% 20.3%	$\begin{array}{c} 13.5\% \ (6.5\%-22.5\%) \\ 21.3\% \ (13\%-31\%) \end{array}$	Q=13, df=4, p=0.011, I ² : 70%, (95%CI: 21%–88%) Q=50.3, df=9, p<0.0001, I ² : 82% (95% CI: 68%–90%)	0.24
$\begin{array}{l} Open \ tib.\\ \leq 48 \ hrs\\ \geq 72 \ hrs \end{array}$	ial fractures 3 (5,26,32) 4 (24,30,37,40)	233 323	$2-22\% \\ 0-15.5\%$	17.5% 9.5%	$13\% (3.6\%-27\%) \\8\% (2\%-17\%)$	Q=13, df=2, p=0.0015, 1 ² : 85%, (95% CI: 54-95%) Q=18, df=3, p=0.0005, 1 ² : 83.2%, (95% CI: 57.4-93.4)	0.38
CI=con ^a random *Mann-	fidence interval; df=degrees of free effects model. Whitney U test.	dom.					

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bacteria including methicillin-resistant *Staphylococcus aureus*. It has a long half-life, low toxicity, and bolus administration is possible [55].

This meta-analysis and review is limited by the lack of high-quality studies specifically looking at the duration of antibiotic administration. The only level 1a study (as per Oxford Centre of Evidence-based Medicine ranking system [56]) included in our analysis was Carsenti-Etasse et al [17]. In addition, only recent publications (5) are clearly stating their definition of infection (Centers for Disease Control and Prevention criteria) [57,58]. Geographical variations and different microbiologic environments have led to a wide spectrum of differing antibiotic regimes making direct comparison of specific antibiotic agents impossible. Treatment standards have only emerged slowly over the decades, and guidelines keep changing, which again makes analysis difficult. The inter-observer reliability of adequately grading and classifying open fractures remains another minefield. The centralization of open fracture treatment to major trauma centers in the United Kingdom will potentially increase the reliability of data.

Conclusions

Therefore, we support the recommendations by BOA/ BAPRAS for open tibial fractures and recommend that they should be extended to all open long bone fractures. Because even shorter than 72 hours regimes did not lead to worse infection-related outcomes, the prolonged antibiotic administration (more than 72 h) should be abandoned. We suggest a randomized trial focusing on different duration of antibiotic treatment for the most severe variants of open fractures or those with gross contamination (fresh water open fractures, farmyard trauma, open pelvic fractures, open intra-articular fractures with exposed joints, high risk hosts (i.e., those with diabetes mellitus, immunosuppressed).

Author Disclosure Statement

No competing financial interests exist.

References

- Hauser CJ, Adams CA, Jr., Eachempati SR. Surgical Infection Society guideline: Prophylactic antibiotic use in open fractures: An evidence-based guideline. Surg Infect 2006;7:379–405.
- Papakostidis C, Kanakaris NK, Pretel J, et al. Prevalence of complications of open tibial shaft fractures stratified as per the Gustilo-Anderson classification. Injury 2011;42:1408– 1415.
- Hoth JJ, Franklin GA, Stassen NA, et al. Prophylactic antibiotics adversely affect nosocomial pneumonia in trauma patients. J Trauma 2003;55:249–254.
- Gosselin RA, Roberts I, Gillespie WJ. Antibiotics for preventing infection in open limb fractures. Cochrane Database Syst Rev 2004;1:CD003764.
- 5. Lack WD, Karunakar MA, Angerame MR, et al. Type III open tibia fractures: Immediate antibiotic prophylaxis minimizes infection. J Orthop Trauma 2015;29:1–6.
- 6. Wilkins J, Patzakis M. Choice and duration of antibiotics in open fractures. Orthop Clin North Am 1991;22:433–437.
- 7. Ritter MA, Campbell E, Keating EM, Faris PM. Comparison of intraoperative versus 24 hour antibiotic prophylaxis

in total joint replacement. A controlled prospective study. Orthop Rev 1989;18:694–696.

- Naique SB, Pearse M, Nanchahal J. Management of severe open tibial fractures: The need for combined orthopaedic and plastic surgical treatment in specialist centres. J Bone Joint Surg Br 2006;88:351–357.
- 9. Penn-Barwell JG, Murray CK, Wenke JC. Early antibiotics and debridement independently reduce infection in an open fracture model. J Bone Joint Surg Br 2012;94:107–112.
- National Institute for Health and Care Excellence (NICE) (2016): Fractures (complex): Assessment and management. NICE guideline [NG37] www.nice.org.uk/guidance/ng37 (Last accessed September 7, 2017).
- 11. Wordsworth M, Lawton G, Nathwani D, et al. Improving the care of patients with severe open fractures of the tibia: The effect of the introduction of Major Trauma Networks and national guidelines. Bone Joint J 2016;98-B:420–424.
- 12. Moher D, Liberati A, Tetzlaff J, et al. Preferred reporting items for systematic reviews and meta-analyses: The PRISMA statement. BMJ 2009;339:b2535.
- Coleman BD, Khan KM, Maffulli N, et al. Studies of surgical outcome after patellar tendinopathy: Clinical significance of methodological deficiencies and guidelines for future studies. Victorian Institute of Sport Tendon Study Group. Scand J Med Sci Sports 2000;10:2–11.
- Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. BMJ 2003;327:557– 560.
- 15. Cochran W. The combination of estimates from different experiments. Biometrics 1954;10:101–129.
- Barton CA, McMillian WD, Crookes BA, et al. Compliance with the Eastern Association for the Surgery of Trauma guidelines for prophylactic antibiotics after open extremity fracture. Int J Crit Illn Inj Sci 2012;2:57–62.
- Carsenti-Etesse H, Doyon F, Desplaces N, et al. Epidemiology of bacterial infection during management of open leg fractures. Eur J Clin Microbiol Infect Dis 1999;18:315– 323.
- Dellinger EP, Caplan ES, Weaver LD, et al. Duration of preventive antibiotic administration for open extremity fractures. Arch Surg 1988;123:333–339.
- 19. Merritt K. Factors increasing the risk of infection in patients with open fractures. J Trauma 1988;28:823–827.
- Rodriguez L, Jung HS, Goulet JA, et al. Evidence-based protocol for prophylactic antibiotics in open fractures: Improved antibiotic stewardship with no increase in infection rates. J Trauma Acute Care Surg 2014;77:400– 407.
- Bednar DA, Parikh J. Effect of time delay from injury to primary management on the incidence of deep infection after open fractures of the lower extremities caused by blunt trauma in adults. J Orthop Trauma 1993;7:532–535.
- Benson DR, Riggins RS, Lawrence RM, et al. Treatment of open fractures: A prospective study. J Trauma 1983;23: 25–30.
- 23. Bergman BR. Antibiotic prophylaxis in open and closed fractures: A controlled clinical trial. Acta Orthop Scand 1982;53:57–62.
- Braun R, Enzler MA, Rittmann WW. A double-blind clinical trial of prophylactic cloxacillin in open fractures. J Orthop Trauma 1987;1:12–17.
- Christensen J, Greiff J, Rosendahl S. Fractures of the shaft of the tibia treated with AO-compression osteosynthesis. Injury 1982;13(4):307–314.

- Clancey GJ, Hansen ST, Jr. Open fractures of the tibia: A review of one hundred and two cases. J Bone Joint Surg Am 1978;60:118–122.
- 27. Cole JD, Ansel LJ, Schwartzberg R. A sequential protocol for management of severe open tibial fractures. Clin Orthop Relat Res 1995:84–103.
- DeLong WG, Jr., Born CT, Wei SY, et al. Aggressive treatment of 119 open fracture wounds. J Trauma 1999; 46:1049–1054.
- Dunkel N, Pittet D, Tovmirzaeva L, et al. Short duration of antibiotic prophylaxis in open fractures does not enhance risk of subsequent infection. Bone Joint J 2013;95-B:831–837.
- Franklin JL, Johnson KD, Hansen ST, Jr. Immediate internal fixation of open ankle fractures. Report of thirtyeight cases treated with a standard protocol. J Bone Joint Surg Am 1984;66:1349–1356.
- 31. Gopal S, Majumder S, Batchelor AG, et al. Fix and flap: The radical orthopaedic and plastic treatment of severe open fractures of the tibia. J Bone Joint Surg Br 2000;82: 959–966.
- 32. Gustilo RB, Mendoza RM, Williams DN. Problems in the management of type III (severe) open fractures: A new classification of type III open fractures. J Trauma 1984;24: 742–746.
- Johnson KD, Bone LB, Scheinberg R. Severe open tibial fractures: A study protocol. J Orthop Trauma 1988;2:175–180.
- Lenarz CJ, Watson JT, Moed BR, et al. Timing of wound closure in open fractures based on cultures obtained after debridement. J Bone Joint Surg Am 2010;92:1921–1926.
- 35. Leonidou A, Kiraly Z, Gality H, et al. The effect of the timing of antibiotics and surgical treatment on infection rates in open long-bone fractures: A 6-year prospective study after a change in policy. Strategies Trauma Limb Reconstr 2014;9:167–171.
- Moehring HD, Gravel C, Chapman MW, Olson SA. Comparison of antibiotic beads and intravenous antibiotics in open fractures. Clin Orthop Relat Res 2000:254–261.
- Ostermann PA, Seligson D, Henry SL. Local antibiotic therapy for severe open fractures. A review of 1085 consecutive cases. J Bone Joint Surg Br 1995;77:93–97.
- Patzakis M, Wilkins J. Factors influencing infection rate in open fracture wounds. Clin Orthop Relat Res 1989;243:36–40.
- Patzakis MJ, Bains RS, Lee J, et al. Prospective, randomized, double-blind study comparing single-agent antibiotic therapy, ciprofloxacin, to combination antibiotic therapy in open fracture wounds. J Orthop Trauma 2000;14:529–533.
- 40. Patzakis MJ, Wilkins J, Moore TM. Use of antibiotics in open tibial fractures. Clin Orthop Relat Res 1983;178:31–35.
- 41. Russell GG, Henderson R, Arnett G. Primary or delayed closure for open tibial fractures. J Bone Joint Surg Br 1990;72;1:125–128.
- Seligson D, Ostermann PA, Henry SL, Wolley T. The management of open fractures associated with arterial injury requiring vascular repair. J Trauma 1994;37:938–940.
- Templeman DC, Gulli B, Tsukayama DT, Gustilo RB. Update on the management of open fractures of the tibial shaft. Clin Orthop Relat Res 1998;350:18–25.
- 44. Vasenius J, Tulikoura I, Vainionpaa S, Rokkanen P. Clindamycin versus cloxacillin in the treatment of 240 open fractures. A randomized prospective study. Ann. Chir Gynaecol 1998;87:224–228.
- 45. Patzakis MJ, Harvey JP, Jr., Ivler D. The role of antibiotics in the management of open fractures. J Bone Joint Surg Am 1974;56:532–541.

- Nayagam S, Graham K, Pearse M, Nanchahal J. Reconstructive surgery in limbs: The case for the orthoplastic approach. Ann Plast Surg 2011;66:6–8.
- Jaeger M, Maier D, Kern WV, Sudkamp NP. Antibiotics in trauma and orthopedic surgery—a primer of evidencebased recommendations. Injury 2006;37(Suppl 2):S74–S80.
- Al-Arabi YB, Nader M, Hamidian-Jahromi AR, Woods DA. The effect of the timing of antibiotics and surgical treatment on infection rates in open long-bone fractures: A 9-year prospective study from a district general hospital. Injury 2007;38:900–905.
- 49. Pollak AN, Jones AL, Castillo RC, et al. The relationship between time to surgical debridement and incidence of infection after open high-energy lower extremity trauma. J Bone Joint Surg Am 2010;92:7–15.
- Crowley DJ, Kanakaris NK, Giannoudis PV. Irrigation of the wounds in open fractures. J Bone Joint Surg Br 2007;89:580–585.
- 51. Crowley DJ, Kanakaris NK, Giannoudis PV. Debridement and wound closure of open fractures: The impact of the time factor on infection rates. Injury 2007;38:879–889.
- 52. Glass GE, Barrett SP, Sanderson F, et al. The microbiological basis for a revised antibiotic regimen in highenergy tibial fractures: Preventing deep infections by nosocomial organisms. J Plast Reconstr Aesthet Sur 2011;64:375–380.
- 53. Appelgren P, Hellstrom I, Weitzberg E, et al. Risk factors for nosocomial intensive care infection: A long-term prospective analysis. Acta Anaesthesiol Scand. 2001;45:710–719.

- 54. Kanakaris NK, Thanasas C, Keramaris N, et al. The efficacy of negative pressure wound therapy in the management of lower extremity trauma: Review of clinical evidence. Injury 2007;38(Suppl 5):S9–S18.
- 55. Periti P, Mini E, Mosconi G. Antimicrobial prophylaxis in orthopaedic surgery: The role of teicoplanin. J Antimicrob Chemother 1998;41:329–340.
- Oxford Centre for Evidence-based Medicine—Levels of Evidence (March 2009)Available at: www.cebm.net/oxfordcentre-evidence-based-medicine-levels-evidence-march-2009/ Accessed September 13, 2017.
- 57. Henriksen NA, Meyhoff CS, Wetterslev J, et al. Clinical relevance of surgical site infection as defined by the criteria of the Centers for Disease Control and Prevention. J Hosp Infect 2010;75:173–177.
- Wilson AP, Gibbons C, Reeves BC, et al. Surgical wound infection as a performance indicator: Agreement of common definitions of wound infection in 4773 patients. BMJ 2004;329:720.

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